

Math 1710 Class 29

Type I/II Errors, Power, 2-Sample Prop., and Examples
Dr. Back

Nov. 4, 2009

Power, Type I/II Errors, α , and β

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Power
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2-Sample CI's
and HT's

Given $H_0 : p = p_0$, there are two ways an HT can report an inaccurate result:

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2-Sample CI's
and HT's

Given $H_0 : p = p_0$, there are two ways an HT can report an inaccurate result:

	H_0 true	H_0 false
Retain H_0	<i>Good</i>	<i>Type II Error</i> probability = β depends on value of p
Reject H_0	<i>Type I Error</i> probability = α	<i>Good</i> probability = $1-\beta$ = <i>power</i>

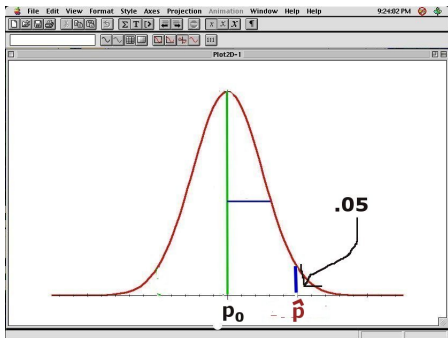
Power, Type I/II Errors, α , and β

Given $H_0 : p = p_0$, there are two ways an HT can report an inaccurate result:

Why $\text{prob}(\text{Type I Error}) = \alpha$?

Borderline value of \hat{p} when $\alpha = .05$

Remember this is the sampling distribution of \hat{p} when $p = p_0$.



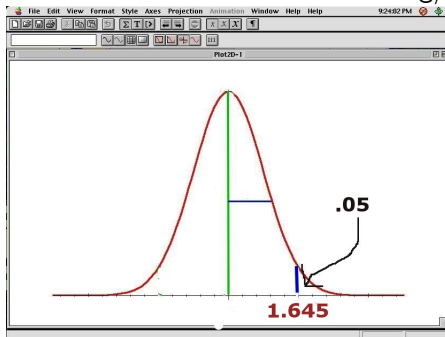
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Given $H_0 : p = p_0$, there are two ways an HT can report an inaccurate result:

Why $\text{prob}(\text{Type I Error}) = \alpha$?

Value of z^* when $\alpha = .05$

This is the borderline value of z in retaining/rejecting.



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Given $H_0 : p = p_0$, there are two ways an HT can report an inaccurate result:

Type I Error Examples:

- a: False Positive in a diagnosis; i.e. deciding a person is sick when they really are not. (H_0 : *The person is well.*)
- b: Convicting an innocent person. (H_0 : *The person is innocent.*)
- c: Producer Risk; the chance that a good good shipment erroneously fails a a test for quality. (H_0 : *A product meets a specification.*)

Power, Type I/II Errors, α , and β

Given $H_0 : p = p_0$, there are two ways an HT can report an inaccurate result:

Type I Error Examples:

- a: False Positive in a diagnosis; i.e. deciding a person is sick when they really are not. (H_0 : *The person is well.*)
- b: Convicting an innocent person. (H_0 : *The person is innocent.*)
- c: Producer Risk; the chance that a good good shipment erroneously fails a a test for quality. (H_0 : *A product meets a specification.*)

Type II Error Examples:

- a: False Negative; missing a sick person.
- b: Letting a guilty person go free.
- c: Consumer Risk; the chance that a bad shipment erroneously passes a a test for quality.

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40% of employees are women.

Woman under-represented as executives?

What would it take in \hat{p} for the company to prove that women are as well represented among executives?

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2-Sample CI's
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40% of employees are women.

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What would it take in \hat{p} for the company to prove that women are as well represented among executives?

$$H_0 : p = .4$$

$$H_a : p > .4$$

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2-Sample CI's
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40% of employees are women.

Woman under-represented as executives?

What would it take in \hat{p} for the company to prove that women are as well represented among executives?

$$H_0 : p = .4$$

$$H_a : p > .4$$

One could also do a HT to see if a given \hat{p} demonstrates women are under represented. Then H_a would be $p < .4$.

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$$H_0 : p = .4$$

$$H_a : p > .4$$

$$n = 420, SD(\hat{p}) = .0239.$$

$$z = \frac{\hat{p} - .4}{.0239}$$

$z > z^*$ means $\hat{p} > .0239z^* + .4$.

$\alpha = .05 \Rightarrow z^* = 1.645$; rejection means $\hat{p} > .439$.

$\alpha = .01 \Rightarrow z^* = 2.326$; rejection means $\hat{p} > .456$.

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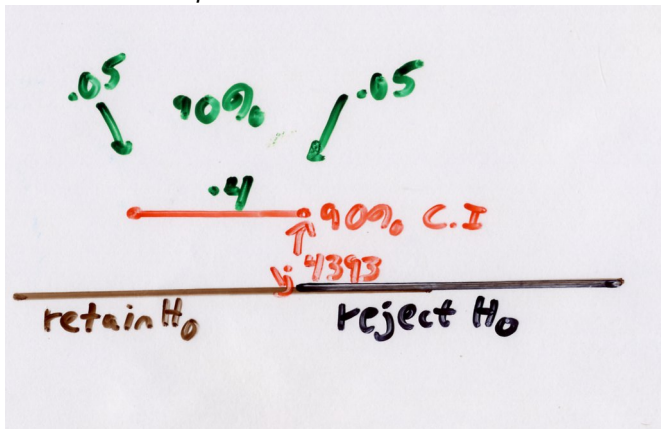
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How a \hat{p} will be dealt with in the HT



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2-Sample CI's
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How does the power depend on the effect size?

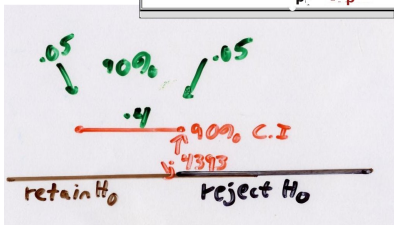
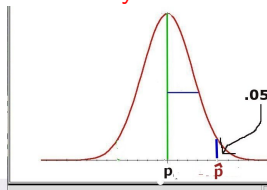
i.e. on the actual value of p

Power Example

How does the power depend on the effect size?

Sampling Dist of \hat{p} when effect size is large.

Power nearly 1

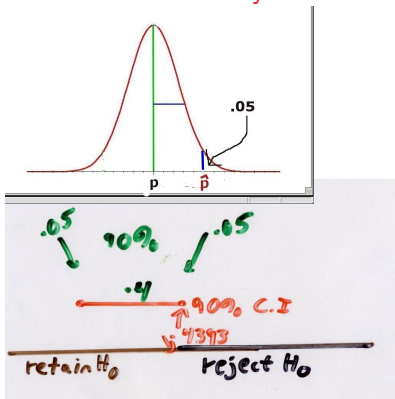


Power Example

How does the power depend on the effect size?

Sampling Dist of \hat{p} when effect size is small.

Power nearly α

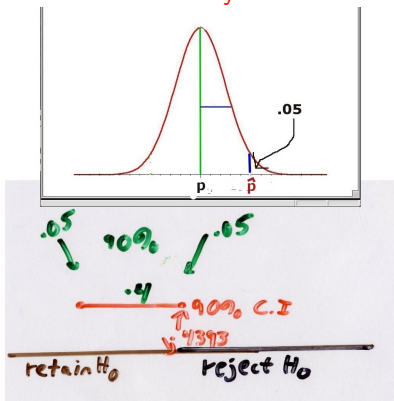


Power Example

How does the power depend on the effect size?

Sampling Dist of \hat{p} when effect size is middle size.

Power in middle between 0 and 1; you could calculate it, but we won't ask you to.



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How does the power depend on the effect size?

$\alpha = .05$ case:

p	β	power
.4001	.95	.05
.41	.89	.11
.45	.33	.67
.5	.01	.99

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How does the power depend on the effect size?

$\alpha = .01$ case:

p	β	power
.4001	.99	.01
.41	.97	.03
.45	.59	.41
.5	.04	.96

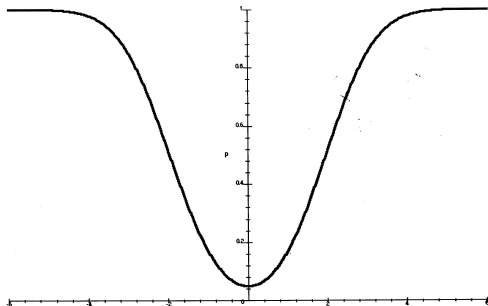
Power Example

How does the power depend on the effect size?

Power vs. alternative value of μ in 2-sided case.

A Power Curve

Power Curve Against $\mu = a$, $\sigma = 1$, H_0 is $\mu = 0$



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Power Properties

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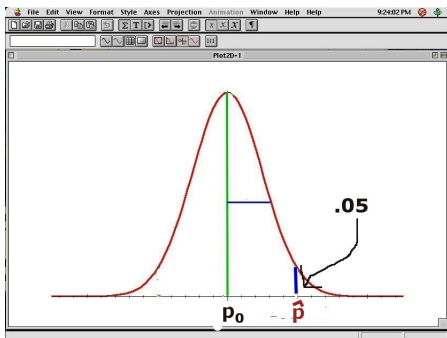
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Prob of type I error = α .

(H_0 1-sides, $\alpha = .05$ below)



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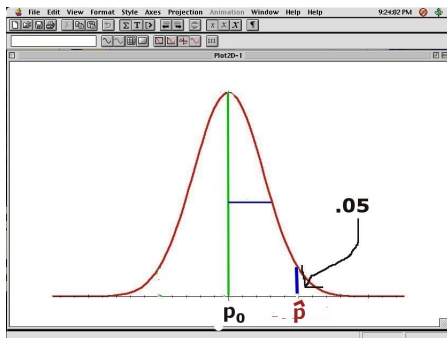
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Power = $1 - \beta$ always.



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Power depends on effect size. (i.e actual alternative value for p .)

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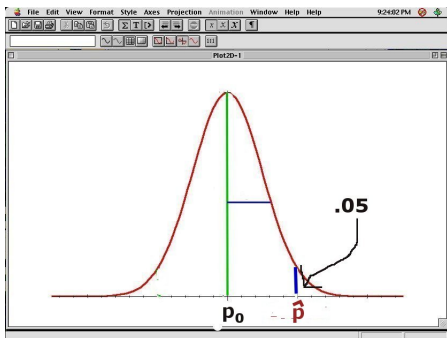
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Power is approx. α when actual p is near p_0 but not exactly p_0 .



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Power goes to 1 as effect size grows, assuming in the 1 sided case that the alternative value supports H_a .

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Power increases as sample size increases.

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2-Sample CI's
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Increasing α decreases β . (Easier to reject H_0 .)

300 stdnts, 60% approve; 200 facilty, 65%

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A significant difference? CI for difference in rates of approval?

300 stdnts, 60% approve; 200 facilty, 65%

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A significant difference? CI for difference in rates of approval?

Let p_1 and p_2 denote the true proportions of students and faculty that approve.

300 stdnts, 60% approve; 200 facilty, 65%

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and HT's

A significant difference? CI for difference in rates of approval?
2-sample inference based on the sampling distribution of
 $\hat{p}_1 - \hat{p}_2$

300 stdnts, 60% approve; 200 facly, 65%

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2-Sample CI's
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A significant difference? CI for difference in rates of approval?
2-sample inference based on the sampling distribution of
 $\hat{p}_1 - \hat{p}_2$

$$\mu = p_1 - p_2$$

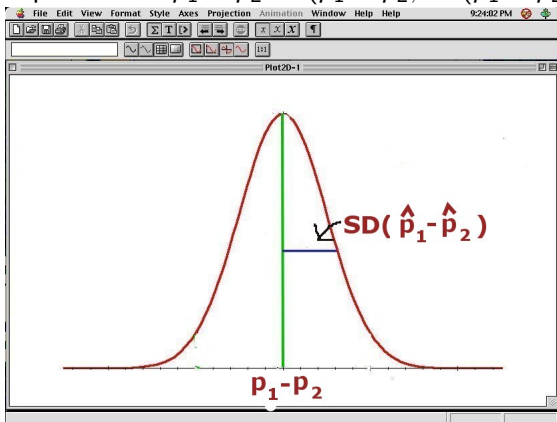
$$\text{Var}(\hat{p}_1 - \hat{p}_2) = \text{Var}(\hat{p}_1) + \text{Var}(\hat{p}_2) = \frac{p_1 q_1}{n_1} + \frac{p_2 q_2}{n_2}$$

$$\text{SD}(\hat{p}_1 - \hat{p}_2) = \sqrt{\frac{p_1 q_1}{n_1} + \frac{p_2 q_2}{n_2}}$$

300 stdnts, 60% approve; 200 facly, 65%

A significant difference? CI for difference in rates of approval?
2-sample inference based on the sampling distribution of $\hat{p}_1 - \hat{p}_2$

Samp. Dist. of $\hat{p}_1 - \hat{p}_2$: $N(p_1 - p_2, SD(\hat{p}_1 - \hat{p}_2))$



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300 stdnts, 60% approve; 200 facilty, 65%

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2-Sample CI's
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2-sample inference based on the sampling distribution of
 $\hat{p}_1 - \hat{p}_2$

Find a CI for $p_1 - p_2$:

300 stdnts, 60% approve; 200 facilty, 65%

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2-Sample CI's
and HT's

2-sample inference based on the sampling distribution of
 $\hat{p}_1 - \hat{p}_2$

Find a CI for $p_1 - p_2$:

Since we don't know p_1 and p_2 , we can't directly compute
 $SD(\hat{p}_1 - \hat{p}_2)$.

So we use $SE(\hat{p}_1 - \hat{p}_2)$ instead.

300 stdnts, 60% approve; 200 facly, 65%

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2-Sample CI's
and HT's

2-sample inference based on the sampling distribution of
 $\hat{p}_1 - \hat{p}_2$

Find a CI for $p_1 - p_2$:

$$SE(\hat{p}_1 - \hat{p}_2) = \sqrt{\frac{\hat{p}_1 \hat{q}_1}{n_1} + \frac{\hat{p}_2 \hat{q}_2}{n_2}}$$

300 stdnts, 60% approve; 200 facilty, 65%

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2-Sample CI's
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2-sample inference based on the sampling distribution of

$$\hat{p}_1 - \hat{p}_2$$

Find a CI for $p_1 - p_2$:

$$SE(\hat{p}_1 - \hat{p}_2) = \sqrt{\frac{\hat{p}_1 \hat{q}_1}{n_1} + \frac{\hat{p}_2 \hat{q}_2}{n_2}}$$

Same argument as in the 1-sample case gives a CI for $p_1 - p_2$ of

$$\hat{p}_1 - \hat{p}_2 \pm z^* SE(\hat{p}_1 - \hat{p}_2).$$

300 stdnts, 60% approve; 200 faculty, 65%

2-sample inference based on the sampling distribution of

$$\hat{p}_1 - \hat{p}_2$$

Find a CI for $p_1 - p_2$:

$$SE(\hat{p}_1 - \hat{p}_2) = \sqrt{\frac{\hat{p}_1 \hat{q}_1}{n_1} + \frac{\hat{p}_2 \hat{q}_2}{n_2}}$$

Same argument as in the 1-sample case gives a CI for $p_1 - p_2$ of

$$\hat{p}_1 - \hat{p}_2 \pm z^* SE(\hat{p}_1 - \hat{p}_2).$$

Here we have

$$SE(\hat{p}_1 - \hat{p}_2) = \sqrt{\frac{.6 \cdot .4}{300} + \frac{.65 \cdot .35}{200}} = .0440.$$

300 stdnts, 60% approve; 200 facly, 65%

2-sample inference based on the sampling distribution of

$$\hat{p}_1 - \hat{p}_2$$

Find a CI for $p_1 - p_2$:

Same argument as in the 1-sample case gives a CI for $p_1 - p_2$ of

$$\hat{p}_1 - \hat{p}_2 \pm z^* SE(\hat{p}_1 - \hat{p}_2).$$

Here we have

$$SE(\hat{p}_1 - \hat{p}_2) = \sqrt{\frac{.6 \cdot .4}{300} + \frac{.65 \cdot .35}{200}} = .0440.$$

A 95% CI for $p_1 - p_2$ is:

$$(.6 - .65) \pm 1.96 \cdot .0440 = -.05 \pm .0863 = (-.1363, .0363).$$

300 stdnts, 60% approve; 200 facilty, 65%

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2-Sample CI's
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Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

300 stdnts, 60% approve; 200 facilty, 65%

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2-Sample CI's
and HT's

Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

Without the request for P-value, we could use the CI above. But for the P-value we need to use "Method 1."

300 stdnts, 60% approve; 200 facilty, 65%

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2-Sample CI's
and HT's

Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

Our hypotheses are:

- $H_0: p_1 = p_2$
- $H_a: p_1 \neq p_2$

300 stdnts, 60% approve; 200 faculty, 65%

Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

Our hypotheses are:

- $H_0: p_1 = p_2$
- $H_a: p_1 \neq p_2$

A twist enters. We are only interested in the reasonableness of our observed $\hat{p}_1 - \hat{p}_2$ with respect to the sampling dist if H_0 is true. There are many such distributions (since we don't know the common value of $p_1 = p_2$ to use.) In particular what we did with $SE(\hat{p}_1 - \hat{p}_2)$ above does not fit the $p_1 = p_2$ situation.

300 stdnts, 60% approve; 200 faculty, 65%

Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

Our hypotheses are:

- $H_0: p_1 = p_2$
- $H_a: p_1 \neq p_2$

We resolve this conflict by making our best estimate of the common value of p_1 and p_2 , namely the weighted average

$$\hat{p}_{pooled} = \frac{n_1 \hat{p}_1 + n_2 \hat{p}_2}{n_1 + n_2}$$

and then

$$SE_{pooled}(\hat{p}_1 - \hat{p}_2) = \sqrt{\frac{\widehat{p}_{pooled} \widehat{q}_{pooled}}{n_1} + \frac{\widehat{p}_{pooled} \widehat{q}_{pooled}}{n_2}}$$

300 stdnts, 60% approve; 200 faculty, 65%

Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

Here the weighted average is

$$\hat{p}_{pooled} = \frac{300 \cdot .60 + 200 \cdot .65}{200 + 300} = .6 \cdot 300 + .4 \cdot .65 = .63$$

and then

$$SE_{pooled}(\hat{p}_1 - \hat{p}_2) = \sqrt{\frac{.63 \cdot .37}{300} + \frac{.63 \cdot .37}{200}} = .0443.$$

300 stdnts, 60% approve; 200 facilty, 65%

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Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

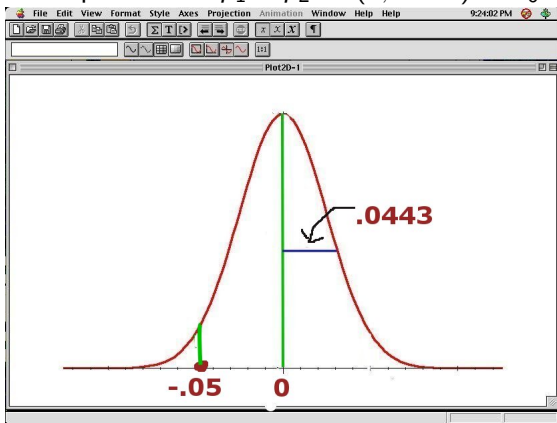
Our z-statistic is

$$z = \frac{\hat{p}_1 - \hat{p}_2}{SE_{pooled}(\hat{p}_1 - \hat{p}_2)} = \frac{-.05}{.0443} = -1.12.$$

300 stdnts, 60% approve; 200 faculty, 65%

Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

Approx Samp. Dist. of $\hat{p}_1 - \hat{p}_2$: $N(0, .0443)$ if H_0 is true.



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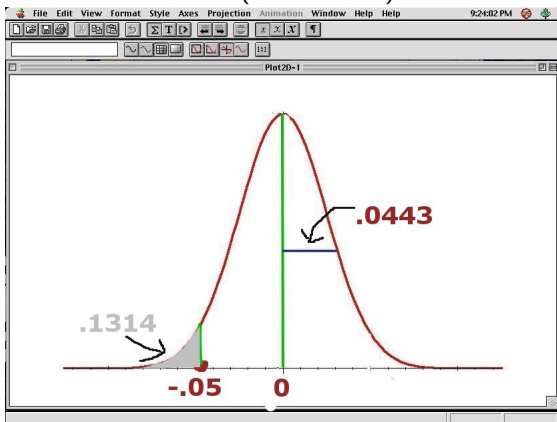
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300 stdnts, 60% approve; 200 faculty, 65%

Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

Tail Prob. is $P(Z < -1.12) = .1314$.



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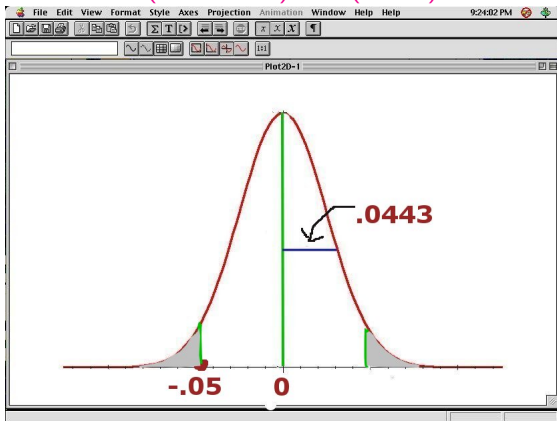
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300 stdnts, 60% approve; 200 faculty, 65%

Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

$$P\text{-value} = 2(\text{Tail Prob.}) = 2(.1314) = .2628$$



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300 stdnts, 60% approve; 200 facilty, 65%

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Carry out an HT at a sig level of $\alpha = .05$ of whether faculty and student approval rates are different. Calculate the P-value as well.

Our P-value is larger than $\alpha = .05$, so we retain H_0 .