

Math 1710 Class 5

Random Variable Operations, Reverse Conditioning Dr. Back

Sep. 7, 2009

Announcements

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Last Time

Operations on
RV's

Behavior of
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Bernoulli and
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Problems

Why the
formulas for
 μ, σ of
 $X + Y$?

Hard
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Problem

Reverse
Conditioning

Please write prominently on your HW which Th Recitation you are *attending*.

Please bring two pennies (or other coins) to class on Wed.

Mean of an RV

The RV X with probability distribution

X	probability
x_1	p_1
x_2	p_2
\dots	\dots
x_n	p_n

Mean:

$$\mu = EX = \sum p_i x_i.$$

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The RV X with probability distribution

X	probability
x_1	p_1
x_2	p_2
\dots	\dots
x_n	p_n

Mean:

$$\mu = EX = \sum p_i x_i.$$

Variance:

$$\text{Var}(X) = \sum p_i (x_i - \mu)^2$$

Standard Deviation:

$$\sigma_X = \sqrt{\text{Var}(X)}.$$

Two Fair Dice

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T	probability
2	$\frac{1}{36}$
3	$\frac{2}{36}$
4	$\frac{3}{36}$
5	$\frac{4}{36}$
6	$\frac{5}{36}$
7	$\frac{6}{36}$
8	$\frac{5}{36}$
9	$\frac{4}{36}$
10	$\frac{3}{36}$
11	$\frac{2}{36}$
12	$\frac{1}{36}$

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$$T = U_1 + U_2$$

U	probability
1	$\frac{1}{6}$
2	$\frac{1}{6}$
3	$\frac{1}{6}$
4	$\frac{1}{6}$
5	$\frac{1}{6}$
6	$\frac{1}{6}$

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U	probability
1	$\frac{1}{6}$
2	$\frac{1}{6}$
3	$\frac{1}{6}$
4	$\frac{1}{6}$
5	$\frac{1}{6}$
6	$\frac{1}{6}$

$$\mu = EU = 3.5$$

$$\begin{aligned} \text{Var}(U) &= \frac{1}{6}(1 - 3.5)^2 + \frac{1}{6}(2 - 3.5)^2 + \frac{1}{6}(3 - 3.5)^2 \\ &\quad + \frac{1}{6}(4 - 3.5)^2 + \frac{1}{6}(5 - 3.5)^2 + \frac{1}{6}(6 - 3.5)^2 \\ &= 2.917. \end{aligned}$$

$$\text{and } \sigma = \sqrt{2.917} = 1.708.$$

Three Important Operations

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- Addition of RV's as in $T = U_1 + U_2$.
- Adding a constant: $Y = X + c$.
- Multiplying by a constant: $Y = cX$.

Thinking about cX and $X + c$

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A simple die game:

1, 2, or 3 gives \$0; 4 or 5 gives \$5; 6 gives \$50.

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X	probability
0	$\frac{1}{2}$
5	$\frac{1}{3}$
50	$\frac{1}{6}$

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X	probability
0	$\frac{1}{2}$
5	$\frac{1}{3}$
50	$\frac{1}{6}$

$$\mu = \frac{5}{3} + \frac{25}{3} = 10; \sigma = 18.03.$$

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$$\mu = \frac{5}{3} + \frac{25}{3} = 10; \sigma = 18.03.$$

Doubling the stakes: $Y = 2X$

Y	probability
0	$\frac{1}{2}$
10	$\frac{1}{3}$
100	$\frac{1}{6}$

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$$\mu = \frac{5}{3} + \frac{25}{3} = 10; \sigma = 18.03.$$

Doubling the stakes: $Y = 2X$

Y	probability
0	$\frac{1}{2}$
10	$\frac{1}{3}$
100	$\frac{1}{6}$

$$\mu = 20; \sigma = 36.06.$$

Thinking about cX and $X + c$

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0	$\frac{1}{2}$
5	$\frac{1}{3}$
50	$\frac{1}{6}$

$$\mu = \frac{5}{3} + \frac{25}{3} = 10; \sigma = 18.03.$$

Including a Payment of 12\$ to Play the Game: $W = X - 12$

W	probability
-12	$\frac{1}{2}$
-7	$\frac{1}{3}$
38	$\frac{1}{6}$

Thinking about cX and $X + c$

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$$\mu = \frac{5}{3} + \frac{25}{3} = 10; \sigma = 18.03.$$

Including a Payment of 12\$ to Play the Game: $W = X - 12$

W	probability
-12	$\frac{1}{2}$
-7	$\frac{1}{3}$
38	$\frac{1}{6}$

$$\mu = -2; \sigma = 18.03.$$

Means are very intuitive!

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$$\mu = \sum p_i x_i$$

$$\mu_{X+Y} = \mu_X + \mu_Y$$

$$\mu_{X+c} = \mu_X + c$$

$$\mu_{cX} = c\mu_X$$

(Handy: $\sum p_i = 1.$)

Means are very intuitive!

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Std. Dev. for cX and $X + c$ are pretty natural.

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$$\text{Var}(X) = \sum p_i (x_i - \mu)^2$$

$$\text{Var}(X + c) = \text{Var}(X).$$

$$\sigma_{X+c} = \sigma_X$$

$$\text{Var}(cX) = c^2 \text{Var}(X).$$

$$\sigma_{cX} = |c| \sigma_X. \text{ So } \sigma_{\mathbf{X}} = \sigma_{-\mathbf{X}}.$$

Std. Dev. for cX and $X + c$ are pretty natural.

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$$\text{Var}(X) = \sum p_i (x_i - \mu)^2$$

$$\text{Var}(X + c) = \text{Var}(X).$$

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Variance of a Sum of *Independent* RV's

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There is no general formula for $\text{Var}(X + Y)$ in terms of $\text{Var}(X)$ and $\text{Var}(Y)$.

Variance of a Sum of *Independent* RV's

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There is no general formula for $\text{Var}(X + Y)$ in terms of $\text{Var}(X)$ and $\text{Var}(Y)$.

But if X and Y are independent RV's:

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$$

Variance of a Sum of *Independent* RV's

There is no general formula for $\text{Var}(X + Y)$ in terms of $\text{Var}(X)$ and $\text{Var}(Y)$.

But if X and Y are independent RV's:

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$$

Very important in Statistics!

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Variance of a Sum of *Independent* RV's

There is no general formula for $\text{Var}(X + Y)$ in terms of $\text{Var}(X)$ and $\text{Var}(Y)$.

But if X and Y are independent RV's:

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$$

Very important in Statistics!

Without independence, this often fails; e.g.

$$\begin{aligned}\text{Var}(X + X) &= \text{Var}(2X) = 2^2\text{Var}(X) = 4\text{Var}(X) \\ &\neq \text{Var}(X) + \text{Var}(X) = 2\text{Var}(X)\end{aligned}$$

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X a Bernoulli(p) RV

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X is the number of heads in *one* toss of a coin,
 p being the probability of a head.

X a Bernoulli(p) RV

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X is the number of heads in *one* toss of a coin,
 p being the probability of a head.

X	probability
0	$q = 1 - p$
1	p

X a Bernoulli(p) RV

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μ , σ ?

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X is the number of heads in *one* toss of a coin,
 p being the probability of a head.

X	probability
0	$q = 1 - p$
1	p

$$\mu = q(0) + p(1) = p.$$

X a Bernoulli(p) RV

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0	$q = 1 - p$
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Var(X) ?

X a Bernoulli(p) RV

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 p being the probability of a head.

X	probability
0	$q = 1 - p$
1	p

$$\mu = q(0) + p(1) = p.$$

$$\mathbf{Var}(X) = q(0 - p)^2 + p(1 - p)^2.$$

X a Bernoulli(p) RV

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X	probability
0	$q = 1 - p$
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$$\mu = q(0) + p(1) = p.$$

$$\begin{aligned}\mathbf{Var}(X) &= q(0 - p)^2 + p(1 - p)^2 \\ &= qp^2 + pq^2 = pq(p + q) = pq.\end{aligned}$$

X a Bernoulli(p) RV

X is the number of heads in *one* toss of a coin,
 p being the probability of a head.

X	probability
0	$q = 1 - p$
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$$\mu = q(0) + p(1) = p.$$

$$\begin{aligned}\mathbf{Var}(X) &= q(0 - p)^2 + p(1 - p)^2 \\ &= qp^2 + pq^2 = pq(p + q) = pq.\end{aligned}$$

$$\sigma = \sqrt{pq}.$$

Y a Binomial(n,p) RV

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Y is the number of heads in n tosses of a coin,
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Y is the number of heads in n tosses of a coin,
 p being the probability of a head.

$$Y = X_1 + X_2 + \dots + X_n$$

where X_i is a Bernoulli(p) RV describing the i 'th toss.

Y a Binomial(n,p) RV

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 p being the probability of a head.

$$Y = X_1 + X_2 + \dots + X_n$$

were X_i is a Bernoulli(p) RV describing the i 'th toss.
The X_i are independent copies of X with

$$\mu_X = p, \mathbf{Var}(X) = pq.$$

Y a Binomial(n,p) RV

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$$Y = X_1 + X_2 + \dots + X_n$$

were X_i is a Bernoulli(p) RV describing the i 'th toss.

So $\mu_Y = p + p + \dots + p = np$.

Y a Binomial(n,p) RV

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So $\mu_Y = p + p + \dots + p = np$.

And $\text{Var}(Y) = pq + pq + \dots + pq = npq$.

Y a Binomial(n,p) RV

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where X_i is a Bernoulli(p) RV describing the i 'th toss.

So $\mu_Y = p + p + \dots + p = np$.

And $\text{Var}(Y) = pq + pq + \dots + pq = npq$.

Making $\sigma_Y = \sqrt{npq}$.

Application

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Joan tosses a coin 10,000 times and sees 5,300 heads.

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Joan tosses a coin 10,000 times and sees 5,300 heads.
Note Binomial(10,000,.5) has $\mu = 5,000$, $\sigma = 50$.

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We'll see later that models based on Binomial RV's almost always fall within 3 std. dev. of the mean.

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5300 is 6 std dev from the mean.

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We'll see later that models based on Binomial RV's almost always fall within 3 std. dev. of the mean.

5300 is 6 std dev from the mean.

Joan concludes her coin is probably not fair.

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Suppose a 2-step process:

Mechanical Checkup: \$50 per hour, 90 min average, 15 min standard deviation.

Appearance Prep: \$6 per hour, 60 min average, 5 min standard deviation.

Find μ and σ for:

- Total Expense.
- Difference in Expense of Phases.

Solution:

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Let M be an RV describing the time for the mechanical part.

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Let M be an RV describing the time for the mechanical part.
Let A be an RV describing the time for the appearance part.

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Let M be an RV describing the time for the mechanical part.
Let A be an RV describing the time for the appearance part.
Let T be an RV describing the total expense.

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Let M be an RV describing the time for the mechanical part.
Let A be an RV describing the time for the appearance part.
Let T be an RV describing the total expense.
We're given (in hours):

	μ	σ
M	$\frac{3}{2}$	$\frac{1}{4}$
A	1	$\frac{1}{12}$

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Let M be an RV describing the time for the mechanical part.
Let A be an RV describing the time for the appearance part.
Let T be an RV describing the total expense.

We're given (in hours):

	μ	σ
M	$\frac{3}{2}$	$\frac{1}{4}$
A	1	$\frac{1}{12}$

$$T = 50M + 6A.$$

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	μ	σ
M	$\frac{3}{2}$	$\frac{1}{4}$
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$$\mu_T = 50\mu_M + 6\mu_A.$$

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We're given (in hours):

	μ	σ
M	$\frac{3}{2}$	$\frac{1}{4}$
A	1	$\frac{1}{12}$

$$T = 50M + 6A.$$

$$\mu_T = 50\mu_M + 6\mu_A.$$

$$\mu_T = 50\mu_M + 6\mu_A = 50\left(\frac{3}{2}\right) + 6(1) = 75 + 6 = 81.$$

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We're given (in hours):

	μ	σ
M	$\frac{3}{2}$	$\frac{1}{4}$
A	1	$\frac{1}{12}$

$$T = 50M + 6A.$$

Assuming independence of M and A ,
 $\text{Var}(T) = 50^2\text{Var}(M) + 6^2\text{Var}(A).$

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$$\text{Var}(T) = 50^2 \text{Var}(M) + 6^2 \text{Var}(A).$$

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M	$\frac{3}{2}$	$\frac{1}{4}$
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Assuming independence of M and A ,

$$\text{Var}(T) = 50^2 \text{Var}(M) + 6^2 \text{Var}(A).$$

$$\text{Var}(T) = 50^2 \left(\frac{1}{4}\right)^2 + 6^2 \left(\frac{1}{12}\right)^2.$$

$$\text{Var}(T) = \left(\frac{25}{2}\right)^2 + \left(\frac{1}{2}\right)^2 = \frac{626}{4} = 156.5.$$

Solution Continued:

We're given (in hours):

	μ	σ
M	$\frac{3}{2}$	$\frac{1}{4}$
A	1	$\frac{1}{12}$

$$T = 50M + 6A.$$

Assuming independence of M and A ,

$$\text{Var}(T) = 50^2 \text{Var}(M) + 6^2 \text{Var}(A).$$

$$\text{Var}(T) = 50^2 \left(\frac{1}{4}\right)^2 + 6^2 \left(\frac{1}{12}\right)^2.$$

$$\text{Var}(T) = \left(\frac{25}{2}\right)^2 + \left(\frac{1}{2}\right)^2 = \frac{626}{4} = 156.5.$$

$$\sigma_T = \sqrt{156.5} = 12.51.$$

Solution Continued:

We're given (in hours):

	μ	σ
M	$\frac{3}{2}$	$\frac{1}{4}$
A	1	$\frac{1}{12}$

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Assuming independence of M and A ,

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Is the independence of M and A natural to assume?

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We're given (in hours):

	μ	σ
M	$\frac{3}{2}$	$\frac{1}{4}$
A	1	$\frac{1}{12}$

$$D = 50M - 6A.$$

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$$\mu_D = 50\mu_M - 6\mu_A = 50\left(\frac{3}{2}\right) - 6(1) = 75 - 6 = 69.$$

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We're given (in hours):

	μ	σ
M	$\frac{3}{2}$	$\frac{1}{4}$
A	1	$\frac{1}{12}$

$$D = 50M - 6A.$$

$$\text{Var}(D) = 50^2 \text{Var}(M) + (-6)^2 \text{Var}(A).$$

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$$\sigma_D = \sqrt{156.5} = 12.51.$$

Same std. dev. for D and T !

Two Independent RV's

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X	probability	Y	probability
x_1	p_1	y_1	q_1
x_2	p_2	y_2	q_2
\dots	\dots	\dots	\dots
x_n	p_n	y_m	q_m

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X	probability	Y	probability
x_1	p_1	y_1	q_1
x_2	p_2	y_2	q_2
\dots	\dots	\dots	\dots
x_n	p_n	y_m	q_m

Independence means $P\{X = x_1 \ \& \ Y = y_1\} = p_1 q_1$.

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X	probability	Y	probability
x_1	p_1	y_1	q_1
x_2	p_2	y_2	q_2
\dots	\dots	\dots	\dots
x_n	p_n	y_m	q_m

Independence means $P\{X = x_1 \ \& \ Y = y_1\} = p_1 q_1$.
In this case $X + Y$ would be $x_1 + y_1$.

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So the probability model for $X + Y$ is

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So the probability model for $X + Y$ is *roughly*

$X + Y$	probability
$x_1 + y_1$	$p_1 q_1$
$x_1 + y_2$	$p_1 q_2$
\dots	\dots
$x_n + y_m$	$p_n q_m$

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So the probability model for $X + Y$ is *roughly*

X + Y	probability
$x_1 + y_1$	$p_1 q_1$
$x_1 + y_2$	$p_1 q_2$
\dots	\dots
$x_n + y_m$	$p_n q_m$

(Why *roughly*: Some value c of $X + Y$ may show up several times in the above table and then more precisely $p_i q_j$ is $P\{X = x_i \ \& \ Y = y_j\}$. And $P(\{X + Y = c\})$ would be the sum of several entries. But computing μ, σ won't be affected by not combining such rows.)

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Assuming we know $E(X+Y)=EX+EY$, then $Var(X + Y)$

$$\begin{aligned} &= \sum_{i,j} p_i q_j (x_i + y_j - \mu_X - \mu_Y)^2 \\ &= \sum_{i,j} p_i q_j (x_i - \mu_X)^2 + \sum_{i,j} p_i q_j (y_j - \mu_Y)^2 \\ &\quad + 2 \sum_{i,j} p_i q_j (x_i - \mu_X)(y_j - \mu_Y) \\ &= (\sum_j q_j) Var(X) + (\sum_i p_i) Var(Y) \\ &\quad + 2 (\sum_i p_i (x_i - \mu_X)) (\sum_j q_j (y_j - \mu_Y)) \\ &= 1 \cdot Var(X) + 1 \cdot Var(Y) + 2 \cdot 0 \cdot 0 \\ &= Var(X) + Var(Y). \end{aligned}$$

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where we used $(a + b)^2 = a^2 + 2ab + b^2$,

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Assuming we know $E(X+Y)=EX+EY$, then $Var(X + Y)$

$$\begin{aligned} &= \sum_{i,j} p_i q_j (x_i + y_j - \mu_X - \mu_Y)^2 \\ &= \sum_{i,j} p_i q_j (x_i - \mu_X)^2 + \sum_{i,j} p_i q_j (y_j - \mu_Y)^2 \\ &\quad + 2 \sum_{i,j} p_i q_j (x_i - \mu_X)(y_j - \mu_Y) \\ &= (\sum_j q_j) Var(X) + (\sum_i p_i) Var(Y) \\ &\quad + 2 (\sum_i p_i (x_i - \mu_X)) (\sum_j q_j (y_j - \mu_Y)) \\ &= 1 \cdot Var(X) + 1 \cdot Var(Y) + 2 \cdot 0 \cdot 0 \\ &= Var(X) + Var(Y). \end{aligned}$$

$$\sum p_i = 1,$$

Variance

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Assuming we know $E(X+Y)=EX+EY$, then $Var(X + Y)$

$$\begin{aligned} &= \sum_{i,j} p_i q_j (x_i + y_j - \mu_X - \mu_Y)^2 \\ &= \sum_{i,j} p_i q_j (x_i - \mu_X)^2 + \sum_{i,j} p_i q_j (y_j - \mu_Y)^2 \\ &\quad + 2 \sum_{i,j} p_i q_j (x_i - \mu_X)(y_j - \mu_Y) \\ &= (\sum_j q_j) Var(X) + (\sum_i p_i) Var(Y) \\ &\quad + 2 (\sum_i p_i (x_i - \mu_X)) (\sum_j q_j (y_j - \mu_Y)) \\ &= 1 \cdot Var(X) + 1 \cdot Var(Y) + 2 \cdot 0 \cdot 0 \\ &= Var(X) + Var(Y). \end{aligned}$$

and $\sum p_i (x_i - \mu_X) = 0$.

Proof of Expectation of Mean Formula

$$E(X + Y)$$

$$\begin{aligned} &= \sum_{i,j} (x_i + y_j) P(X = x_i \text{ and } Y = y_j) \\ &= \sum_{i,j} x_i P(X = x_i | Y = y_j) P(Y = y_j) \\ &\quad + \sum_{i,j} y_j P(Y = y_j | X = x_i) P(X = x_i) \\ &= \sum_i x_i (\sum_j P(X = x_i | Y = y_j) P(Y = y_j)) \\ &\quad + \sum_j y_j (\sum_i P(Y = y_j | X = x_i) P(X = x_i)) \\ &= \sum_i x_i P(X = x_i) \\ &\quad + \sum_j y_j P(Y = y_j) \\ &= E(X) + E(Y). \end{aligned}$$

Unsurprising, but not obvious looking.

Proof of Expectation of Mean Formula

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Unsurprising, but not obvious looking.

Did not require X and Y to be independent!

7 Balls Randomly into 5 boxes

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What is the expected number of empty boxes?

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What is the probability that the first box remains empty?

What is the expected number of empty boxes?

The 1st question is a great hint for easily doing the 2nd!

Triple blood test for Down Syndrome

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Suppose

- 1 in 800 pregnant women affected.
- $\frac{8}{9}$ of affected women identified by the test
- $\frac{1}{4}$ of unaffected women show up as false positives.

Triple blood test for Down Syndrome

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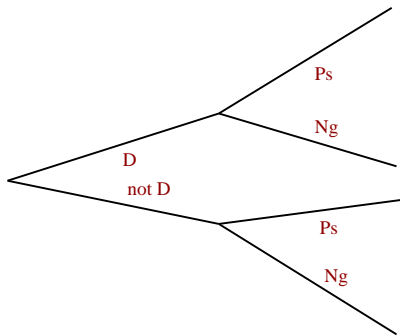
Suppose

- 1 in 800 pregnant women affected.
- $\frac{8}{9}$ of affected women identified by the test
- $\frac{1}{4}$ of unaffected women show up as false positives.

P(Down syndrome if test is positive)?

Notation

D: foetus affected by Downs.
Ps: positive test result reports Downs.
Ng: not Ps. (Negative)



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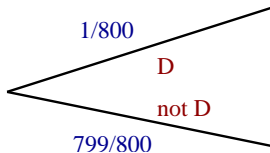
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$$P(D) = \frac{1}{800}$$



With the Conditional Probabilities

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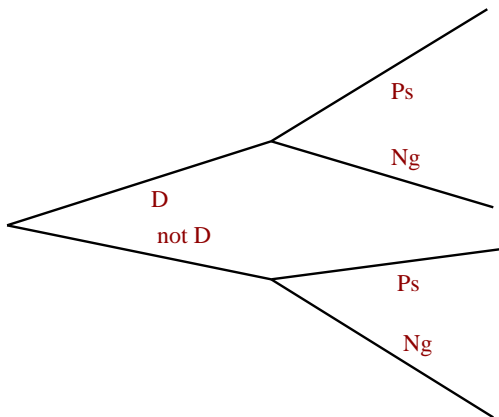
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$$P(Ps|D) = \frac{8}{9},$$

$$P(Ps|\text{not } D) = \frac{1}{4}$$

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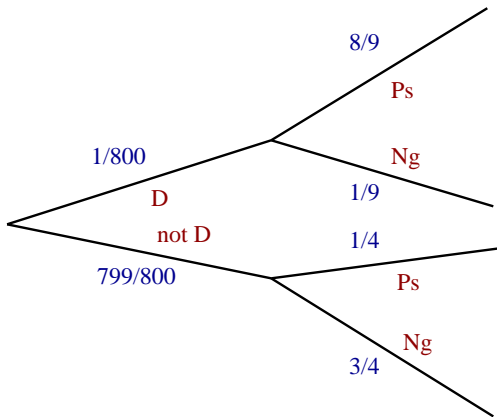
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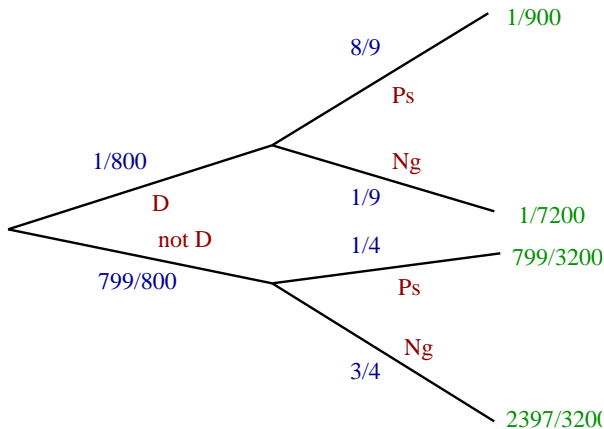
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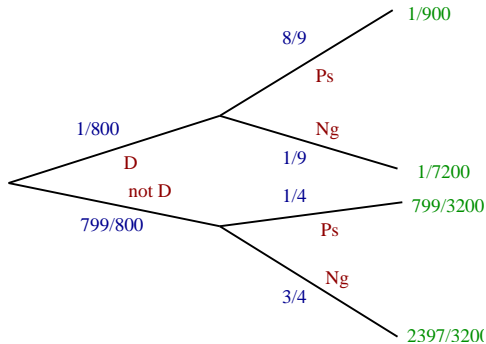
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$$P(D|P_s) = \frac{P(D \text{ and } P_s)}{P(P_s)} = \frac{\frac{1}{900}}{\frac{1}{900} + \frac{799}{3200}} \sim .0044.$$

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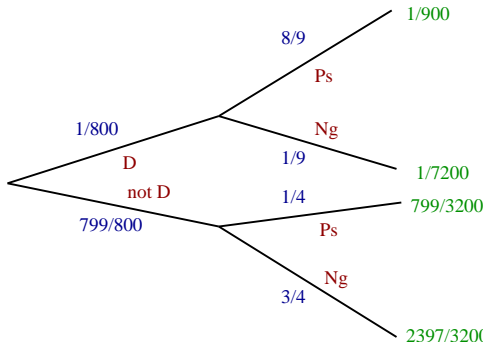
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Not very helpful!