

Linear functions on the compact
classical groups

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Haar measure on O_n

Constructions

1. Fill an $n \times n$ array with independent standard normal random variables, and perform the Gram-Schmidt process.
2. Pick a point at random from S^{n-1} and fill the first column with its coordinates. Fill the second column with a random point in the orthogonal complement (in S^{n-1}) of the first column, and so on.

Meta-theorem: Random orthogonal matrices
are "like" Gaussian matrices.

Some actual theorems along these lines:

Theorem (Foret, 1966): Let X_1 be the first
coordinate of a randomly chosen point on
the sphere S^{n-1} . Then

$$P(\sqrt{n} X_1 \leq t) \xrightarrow{n \rightarrow \infty} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^t e^{-x^2/2} dx$$

i.e., $\sqrt{n} X_1$ is asymptotically Gaussian.

* By construction 2. of Haar measure and its
translation invariance, this means that if
 $M = (m_{ij}) \in O_n$ is a Haar-distributed matrix,
then

$$P(\sqrt{n} m_{k\ell} \leq t) \xrightarrow{n \rightarrow \infty} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^t e^{-x^2/2} dx$$

for each k, ℓ .

A (much) later strengthening:

Theorem (Tomas Blum, 1994):

Let X_i be as before and $Z \sim N(0, 1)$.

Then

$$d_{T.V.}(\sqrt{n} X_i, Z) \leq \frac{4}{n-1}$$

and the bound is sharp up to the value of the constant.

Theorem (Tiang, 2004):

Let $M \sim \text{Haar}(\mathbb{O}_n)$, $k = o(\sqrt{n})$ and Γ_k a $k \times k$ Gaussian matrix. Then if M_k is the top-left $k \times k$ block of M ,

$$d_{TV}(M_k, \Gamma_k) \rightarrow 0 \text{ as } n \rightarrow \infty.$$

Theorem (Tiang, 2004):

For each $n \geq 2$ there are random matrices $M = (m_{ij})$ and $\Gamma = (\gamma_{ij})$ defined on the same probability space st.

1. $M \sim \text{Haar}(\mathbb{O}_n)$, Γ is Gaussian

2. If $\varepsilon_n(m) = \max_{\substack{1 \leq i \leq n \\ 1 \leq j \leq m}} |\sqrt{n} m_{ij} - \gamma_{ij}|$, then if $m = o(\frac{n}{\log n})$,

$\varepsilon_n(m) \rightarrow 0$ in probability as $n \rightarrow \infty$.

Theorem (d'Aristofile / Diaconis / Kiferman, 2003):

Let M be a random $n \times n$ orthogonal matrix, and let $\{k_n\}$ be a sequence of numbers with $k_n \xrightarrow[n \rightarrow \infty]{} \infty$.

Let $\beta_1, \dots, \beta_{k_n}$ be a collection of k_n of the entries of M .

Write

$$S_j^n = \frac{1}{\sqrt{k_n}} \sum_{i=1}^{k_n} \beta_i, \quad W_n(t) = S_{\lfloor k_n t \rfloor}^n$$

Then

$$W_n \xrightarrow[n \rightarrow \infty]{} W, \text{ a standard Brownian motion.}$$

Theorem (M., 2005): Let $M \in \mathcal{O}_n$ be distributed according to Haar measure and $A \in M_n(\mathbb{R})$ fixed, with $\text{tr}(AA^t) = n$. Let $W = \text{tr}(AM)$ and $z \sim N(0,1)$. Then

$$d_{\text{T.V.}}(W, z) \leq \frac{2\sqrt{3}}{n-1}$$

and the rate is sharp up to the constant.

Background for the proof: Stein's method for normal approximation

Idea: The normal distribution is characterized by the differential operator

$$f \xrightarrow{T} f'(x) - xf(x)$$

in the sense that:

1. $\mathbb{E}[f'(z) - zf(z)] = 0$ for $z \sim N(0,1)$.

2. If Y is a random variable s.t.

$$\mathbb{E}[f'(Y) - Yf(Y)] = 0$$

then $Y \sim N(0,1)$.

This follows because the differential equation

$$f'(x) - xf(x) = g(x) - \mathbb{E}g(z)$$

can be solved for f in terms of g :

$$f(x) = U_0 g(x) = e^{x^2/2} \int_{-\infty}^x [g(t) - \mathbb{E}g(z)] e^{-t^2/2} dt.$$

\Rightarrow given g , if $f = U_0 g$, then

$$0 = \mathbb{E}[f'(Y) - Yf(Y)] = \mathbb{E}g(Y) - \mathbb{E}g(z).$$

Next idea:

if $E[f'(Y) - Yf(Y)]$ is small for a large class of f ,
then $Y \sim N(0,1)$.

Fix a test function g , and let $f = U \circ g$.

$$Eg(Y) - Eg(Z) = E[f'(Y) - Yf(Y)] \leftarrow \text{small}$$

Many notions of distance between random variables
(or their distributions) can be expressed as

$$d_{\mathcal{F}}(X, Y) = \sup_{f \in \mathcal{F}} |Ef(X) - Ef(Y)|$$

for some class \mathcal{F} . For example:

$\mathcal{F}_1 = \{ \text{continuous functions vanishing at } \infty \text{ and bounded by } 1 \} \iff \text{total variation distance}$

$\mathcal{F}_2 = \{ f \in C^1(\mathbb{R}) : \|f\|_{\infty} \leq 1, \|f'\|_{\infty} \leq 1 \} \iff \text{dual-Lipschitz distance}$

$\mathcal{F}_3 = \{ f \in C^1(\mathbb{R}) : \|f'\|_{\infty} \leq 1 \} \iff \text{Wasserstein distance}$

$\mathcal{F}_4 = \{ \mathbb{I}_{(-\infty, x]} : x \in \mathbb{R} \} \iff \text{distance between distribution functions}$

Exchangeable pairs

Have: W , conjectured to be normal (approximately).

From W , make a small change to get W' such that (W, W') is exchangeable; i.e., $(W, W') \sim (W', W)$.

Suppose that there is a number λ s.t.

$$\mathbb{E}[W' - W | W] = -\lambda W$$

$$\mathbb{E}[(W' - W)^2 | W] = 2\lambda$$

Then:

$$0 = \mathbb{E}[(W' - W)(f(W') + f(W))]$$

$$= \mathbb{E}[(W' - W)(f(W') - f(W)) + 2(W' - W)f(W)]$$

$$= \mathbb{E}[(W' - W)^2 f'(W) + 2(W' - W)f(W) + R]$$

$$= \mathbb{E}[\mathbb{E}[(W' - W)^2 | W] f'(W) + 2 \mathbb{E}[W' - W | W] f(W) + R]$$

$$= \mathbb{E}[2\lambda f'(W) + 2(-\lambda W)f(W) + R]$$

$$\Rightarrow \mathbb{E}[f'(W) - Wf(W)] = \mathbb{E}\left[-\frac{R}{2\lambda}\right].$$

Theorem: $M \sim \text{Haar}(\mathcal{O}_n)$, $A \in M_n(\mathbb{R})$, $\text{tr}(AA^t) = n$, $W := \text{tr}(AM)$.

$$\Rightarrow d_{TV}(W, Z) \leq \frac{2\sqrt{3}}{n-1}$$

Proof:

Define a family of exchangeable pairs (W, W_ε) as follows.

Define

$$A_\varepsilon = \begin{bmatrix} \sqrt{1-\varepsilon^2} & \varepsilon \\ -\varepsilon & \sqrt{1-\varepsilon^2} \end{bmatrix} \oplus I_{n-2}$$
$$= I_n + \left[\left(-\frac{\varepsilon^2}{2} + O(\varepsilon^4) \right) I_2 + \varepsilon C_2 \right] \oplus O_{n-2}$$

Let $U \sim \text{Haar}(\mathcal{O}_n)$, independent of M .

$$C_2 = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$

then

$$M_\varepsilon := U A_\varepsilon U^t M$$

M_ε is a rotation of M by $\arcsin(\varepsilon)$ in a random 2-dimensional subspace.

Let $W_\varepsilon = \text{tr}(AM_\varepsilon)$; (W, W_ε) is exchangeable for each fixed ε .

As before, fix g and let $f = U \circ g$.

$$0 = \mathbb{E}[(W_\varepsilon - W)(f(W_\varepsilon) + f(W))]$$

$$= \mathbb{E}[(W_\varepsilon - W)^2 f'(W) + 2(W_\varepsilon - W)f(W) + R]$$

$$M_\varepsilon - M = U \left[\left(-\frac{\varepsilon^2}{2} + o(\varepsilon^4) \right) \mathbb{I}_2 + \varepsilon C_2 \right] U^\top M$$

$$= \left(-\frac{\varepsilon^2}{2} + o(\varepsilon^4) \right) K K^\top M + \varepsilon K C_2 K^\top M$$

$$K = \begin{bmatrix} u_{11} & u_{12} \\ \vdots & \vdots \\ u_{n1} & u_{n2} \end{bmatrix}$$

$$C_2 = \begin{bmatrix} 0 & 1 \\ -1 & 0 \end{bmatrix}$$

$$\Rightarrow W_\varepsilon - W = \text{tr}(A(M_\varepsilon - M))$$

$$= \left(-\frac{\varepsilon^2}{2} + o(\varepsilon^4) \right) \text{tr}(A K K^\top M) + \varepsilon \text{tr}(A K C_2 K^\top M)$$

Now, $(K K^\top)_{ij} = u_{i1} u_{j1} + u_{i2} u_{j2} \Rightarrow \mathbb{E}[(K K^\top)_{ij}] = \frac{2}{n} \delta_{ij}$

$$(K C_2 K^\top)_{ij} = u_{i1} u_{j2} - u_{i2} u_{j1} \Rightarrow \mathbb{E}[(K C_2 K^\top)_{ij}] = 0.$$

So $\mathbb{E}[W_\varepsilon - W | W]$

$$= \mathbb{E} \left[\left(-\frac{\varepsilon^2}{2} + o(\varepsilon^4) \right) \mathbb{E}[\text{tr}(A K K^\top M) | M] + \varepsilon \mathbb{E}[\text{tr}(A K C_2 K^\top M) | M] \right]$$

$$= \left(-\frac{\varepsilon^2}{2} + o(\varepsilon^4) \right) \mathbb{E}[\text{tr}(A M) | W]$$

$$= \left(-\frac{\varepsilon^2}{2} + o(\varepsilon^4) \right) W.$$

We had:

$$0 = \mathbb{E} \left[\mathbb{E} [(W_\varepsilon - W)^2 | W] f'(W) + 2 \mathbb{E} [W_\varepsilon - W | W] f(W) + R \right]$$

$$\Rightarrow 0 = \lim_{\varepsilon \rightarrow 0} \frac{1}{\varepsilon^2} \mathbb{E} \left[\mathbb{E} [(W_\varepsilon - W)^2 | W] f'(W) + 2 \mathbb{E} [W_\varepsilon - W | W] f(W) + R \right] \quad *$$

(provided this makes sense)

from the last page,

$$\frac{1}{\varepsilon^2} \mathbb{E} [W_\varepsilon - W | W] = -\frac{1}{n} W + O(\varepsilon^2)$$

By Taylor's theorem,

$$|R| \leq \|f''\|_\infty |W_\varepsilon - W|^3 = O(\varepsilon^3)$$

$$\Rightarrow \frac{1}{\varepsilon^2} \mathbb{E} |R| \rightarrow 0 \text{ as } \varepsilon \rightarrow 0.$$

$$\text{From } W_\varepsilon - W = \left(-\frac{\varepsilon^2}{2} + O(\varepsilon^4) \right) + \text{tr}(AK(K^*M)) + \varepsilon \text{tr}(AKC_2K^*M),$$

$$\mathbb{E} [(W_\varepsilon - W)^2 | W] = \varepsilon^2 \mathbb{E} [\text{tr}(AKC_2K^*M)^2 | W] + O(\varepsilon^3).$$

$$\Rightarrow \frac{1}{\varepsilon^2} \mathbb{E} [(W_\varepsilon - W)^2 | W] = \frac{2}{n} - \frac{2}{n(n-1)} \mathbb{E} [\text{tr}(AM)^2 | W] + O(\varepsilon).$$

$$\Rightarrow \text{by } * \quad 0 = \mathbb{E} \left[\frac{2}{n} f'(W) - \frac{2}{n} W f(W) - \frac{2}{n(n-1)} \mathbb{E} [\text{tr}(AM)^2 | W] f'(W) \right]$$

$$\Rightarrow \mathbb{E} [f'(W) - W f(W)] = \frac{1}{n-1} \underbrace{\mathbb{E} [\text{tr}(AM)^2 | W]}_{\text{bounded}} \underbrace{f'(W)}_{\text{bounded for } g \text{ bounded}}$$

\parallel
 $\mathbb{E} g(W) - \mathbb{E} g(z)$

Things to take away from the proof

1. Stein's method is clever and powerful.
2. Taking advantage of the continuous symmetries of \mathbb{C}^n led to a major improvement (it allowed the derivative approximation to become exact in the limit.)
3. The argument is quite general and doesn't use that much about Haar measure, at least to get started.

Other directions

1. U_n : If a random unitary matrix,
 $A \in M_n(\mathbb{C})$, $\text{tr}(AA^*) = n$, $W := \text{tr}(AM)$.

W is approximately normal but this
is a multivariate problem, so the approach
has to be modified.

2. Multivariate versions for $\mathbb{Q}_n, \mathbb{R}_n$

We can do it, but only for the natural
dual-Lipschitz metric on 'ave' sets!

metrizes for var^* steps!

Exercise 1.1 (Lecture 1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 14, 15, 16, 17, 18, 19, 20, 21, 22, 23, 24, 25, 26, 27, 28, 29, 30, 31, 32, 33, 34, 35, 36, 37, 38, 39, 40, 41, 42, 43, 44, 45, 46, 47, 48, 49, 50, 51, 52, 53, 54, 55, 56, 57, 58, 59, 60, 61, 62, 63, 64, 65, 66, 67, 68, 69, 70, 71, 72, 73, 74, 75, 76, 77, 78, 79, 80, 81, 82, 83, 84, 85, 86, 87, 88, 89, 90, 91, 92, 93, 94, 95, 96, 97, 98, 99, 100):

Let M be a random orthogonal matrix and let A_1, \dots, A_k be $n \times n$ matrices (fixed) with $\text{tr}(A_i A_j^t) = n \delta_{ij}$. Consider the random vector

$$W = (\text{tr}(A_1 M), \dots, \text{tr}(A_k M)).$$

If $Z \in \mathbb{R}^k$ is a standard Gaussian random vector, then

$$d_{L^k}(W, Z) := \sup_{\substack{f, \\ \|f\|_\infty + \|f'\|_\infty \leq 1}} |\mathbb{E}f(W) - \mathbb{E}f(Z)| \leq \frac{C k^{3/2}}{n}$$

for an explicit constant C independent of k and n .

A very similar proof to the proof of the orthogonal result gives:

Theorem (N., 2005):

Let \mathbb{X} be a compact locally symmetric space, and let $f: \mathbb{X} \rightarrow \mathbb{R}$ be an eigenfunction, ^{with non-zero} of ^{eigenvalue.} the Laplacian Δ on \mathbb{X} . Let X be a random point of \mathbb{X} , and consider the random variable $W = f(X)$. If $Z \sim N(0, 1)$, the following bound holds

$$d_{TV}(W, Z) \leq \frac{1}{\lambda} \sqrt{\text{Var}(\|\nabla f\|^2)}.$$

(convention: $\lambda > 0$.)