

MATH 4740 HW8 Solution.

textbook Exercises:

2.18. To find $\mathbb{P}(X=1)$, suppose:

①. $X \sim \text{Binomial}(n=20, p=0.1)$

$$\Rightarrow \mathbb{P}(X=1) = \binom{20}{1} 0.1^1 \cdot 0.9^{20-1} = \underline{0.27017}.$$

②. $X \sim \text{Poisson}(\lambda = np = 20 \cdot 0.1)$

$$\Rightarrow \mathbb{P}(X=1) = e^{-2} \cdot \frac{2^1}{1!} = \underline{0.27067}$$

Comparing this result to the one from ①, we see that Poisson is not a bad approximation to Binomial.

□

2.22

We know that $N(t) \sim \text{Poisson}(3t)$

$T_n \sim \text{Gamma}(n, 3)$, $t_n \sim \text{exponential}(3)$

(a). $\mathbb{E}(T_{12}) = \frac{12}{3} = 4.$

(b). By the memoryless property of t_n ,

$$\mathbb{E}[T_{12} \mid N(2) = 5] = \mathbb{E}\left[\sum_{i=1}^{12} t_i \mid N(2) = 5\right]$$

$$= \mathbb{E}\left[\sum_{i=1}^{12-5} t_i\right] + 2 \quad \left(\text{technically, } \mathbb{E}\left[\sum_{i=6}^{12} t_i\right] + 2\right)$$

$$= 7 \cdot \frac{1}{3} + 2$$

$$= \frac{13}{3}$$

$$\mathbb{E}[N(5) \mid N(2) = 5]$$

$$= \mathbb{E}[N(5) - N(2) + N(2) - N(0) \mid N(2) - N(0) = 5]$$

$$= \mathbb{E}[N(5) - N(2)] + \mathbb{E}[N(2) - N(0) \mid N(2) - N(0) = 5] \quad (\text{independent increments})$$

$$= \mathbb{E}[N(5)] - \mathbb{E}[N(2)] + 5$$

$$= 3.5 - 2.3 + 5$$

$$= 14$$

□

2.34

Denote the weight of i -th trout by Y_i , the number of trouts caught by time t by $N(t)$. Then

$$N(t) \sim \text{Poisson}(3t).$$

In 2 hours, Edwin is expected to catch $\mathbb{E}(N(2)) = 6$ trouts.

$$\text{So } \mathbb{E}\left(\sum_{i=1}^{N(2)} Y_i\right) = \mathbb{E}[N(2)] \mathbb{E}[Y_i]$$

$$= 6 \cdot 4.$$

$$= 24$$

$$\text{Var}\left(\sum_{i=1}^{N(2)} Y_i\right) = \mathbb{E}[N(2)] \text{var}(Y_i) + \text{Var}(N(2)) \mathbb{E}(Y_i)^2$$

$$= 6 \cdot 2^2 + 6 \cdot 4^2$$

$$= 120.$$

$$\text{sd}\left(\sum_{i=1}^{N(2)} Y_i\right) = \sqrt{120} = 2\sqrt{30}.$$

□

2.42

Denote the number of trucks passed by time t as $N_T(t)$

the number of cars passed by time t as $N_C(t)$.

$$N_T(t) \sim \text{Poisson} \left(\frac{2}{3} \cdot 0.1 \cdot t \right)$$

$$N_C(t) \sim \text{Poisson} \left(\frac{2}{3} \cdot 0.9 \cdot t \right)$$

$$N_C(t) \perp\!\!\!\perp N_T(t).$$

$$(a) \quad \mathbb{P} (N_T(60) \geq 1)$$

$$= 1 - \mathbb{P} (N_T(60) = 0)$$

$$= 1 - e^{-4} \cdot \frac{4^0}{0!}$$

$$= 1 - e^{-4}$$

$$(b) \quad \mathbb{E} [N_T(60) + N_C(60) \mid N_T(60) = 10]$$

$$= \mathbb{E} [N_C(60)] + 10$$

$$= \frac{2}{3} \cdot 0.9 \cdot 60 + 10$$

$$= 46.$$

$$(c) \quad \mathbb{P} (N_T(60) = 5, N_C(60) = 45 \mid N_T(60) + N_C(60) = 50)$$

$$= \binom{50}{5} 0.1^5 \cdot 0.9^{45}$$

$$= 0.1849$$

□

2.57.

$$\begin{aligned} (a) \quad & \mathbb{P}(N(3)=4 \mid N(1)=1) \\ &= \mathbb{P}(N(3) - N(1) + N(1) - N(0) = 4 \mid N(1) - N(0) = 1) \\ &= \mathbb{P}(N(3) - N(1) = 4 - 1) \\ &= \mathbb{P}(N(2) = 3) \\ &= e^{-2 \cdot 2} \cdot \frac{(2 \cdot 2)^3}{3!} \\ &= 0.195. \end{aligned}$$

$$\begin{aligned} (b) \quad & \mathbb{P}(N(1)=1 \mid N(3)=4) \\ &= \binom{4}{1} \left(\frac{1}{3}\right)^1 \left(\frac{2}{3}\right)^3 \\ &= 0.395. \end{aligned}$$

□

Additional Problems:

1. A Poisson Process $\{N(s)\}$ with rate λ is a stochastic process constructed by a sequence of independent, continuously - occurring events where $N(t) \sim \text{Poisson}(\lambda t)$ and $N(t)$'s are related via $\textcircled{1} N(t) - N(s) \sim \text{Poisson}(\lambda(t-s))$ for all $t > s$ and $\textcircled{2}$ independent increments; whilst a Poisson random variable $N(s)$ is just a random variable, the value of which follows a probability measure as in Poisson distribution,

For any point in the 4-dimensional cube of side length t ,
2. \checkmark , the ordering of (t_1, t_2, t_3, t_4) has $4! = 24$ possibilities
(a)

with equal probability, for example

$$\mathbb{P}(t_1 < t_2 < t_3 < t_4) = \mathbb{P}(t_2 < t_1 < t_4 < t_3) = \dots = \frac{1}{24}$$

So the space $\{(t_1, t_2, t_3, t_4) \in \mathbb{R}^4 : 0 \leq t_1 \leq t_2 \leq t_3 \leq t_4 \leq t\}$

has the same volume as $\{(t_1, t_2, t_3, t_4) \in \mathbb{R}^4 : 0 \leq t_2 \leq t_1 \leq t_4 \leq t_3 \leq t\}$
and all other 22 subspaces.

By symmetry, the volume of any such subspace is $\frac{t^4}{24}$.

(b) If the given joint density is correct, then

$$\begin{aligned} & \int_A f(t_1, t_2, t_3, t_4) dt_1 dt_2 dt_3 dt_4 \\ &= \int_A \lambda^4 e^{-\lambda t} dt_1 dt_2 dt_3 dt_4 \\ &= \lambda^4 e^{-\lambda t} \cdot \text{vol}(A) \\ &= e^{-\lambda t} \frac{(\lambda t)^4}{4!} \end{aligned}$$

However, the integral should be equal to 1 since it's integrating over a density function. The problem with the given joint density is that it needs to be divided by $\mathbb{P}(N(t)=4)$ because we are conditioning the probability on $N(t)=4$. So the correct joint

density is given by: $f(t_1, t_2, t_3, t_4) = \frac{\lambda^4 e^{-\lambda t}}{e^{-\lambda t} \cdot \frac{(\lambda t)^4}{4!}} = \frac{4!}{t^4}$

(c) Since $U_i \sim \text{uniform}[0, t]$, we can view (U_1, U_2, U_3, U_4) as a random point within the 4-dimensional cube of side length t as described in (a). Therefore the joint density would be $\frac{1}{\text{vol}(A)} = \frac{4!}{t^4}$

And we know that the joint density of (T_1, T_2, T_3, T_4) conditioned on $N(t) = 4$ is given by $\frac{4!}{t^4}$ as in part (b), which is the same density function of V_i 's.

□

3.(a)

(i) $M(0) = N(t_0) - N(t_0) = 0$

(ii)
$$\begin{aligned} M(t+s) - M(s) &= (N(t_0) - N(t_0 - t - s)) - (N(t_0) - N(t_0 - s)) \\ &= N(t_0 - s) - N(t_0 - t - s) \\ &= N(t_0 - s - t + t) - N(t_0 - t - s) \\ &\sim \text{Poisson}(\lambda t) \end{aligned}$$

(iii) consider $0 \leq s_1 \leq s_2 \leq \dots \leq s_n \leq t_0$

$$\begin{aligned} M(s_{i+1}) - M(s_i) &= N(t_0) - N(t_0 - s_{i+1}) - N(t_0) + N(t_0 - s_i) \\ &= N(t_0 - s_i) - N(t_0 - s_{i+1}) \end{aligned}$$

since $0 \leq t_0 - s_n \leq t_0 - s_{n-1} \leq \dots \leq t_0 - s_2 \leq t_0 - s_1 \leq t_0$,

we have $N(t_0 - s_i) - N(t_0 - s_{i+1})$ are independent for $i=1, \dots, n$ and conclude $M(s)$ has independent increments too.

(b). By (a) we know that $\{M(s)\}$ is a backward Poisson process,

so $t-L$ is just the first arrival of $M(s)$. Also notice that $L=0$

if there is no arrival by time t , which occurs with probability

$1 - \int_0^t x\lambda e^{-\lambda x} dx = e^{-\lambda t}$. Therefore,

$$\mathbb{E}(t-L) = \int_0^t x\lambda e^{-\lambda x} dx + (t-0) \cdot e^{-\lambda t}$$

$$= \frac{1}{\lambda} (1 - e^{-\lambda t})$$

As $t \rightarrow \infty$, $e^{-\lambda t} \rightarrow 0$, $\mathbb{E}(t-L) \rightarrow \frac{1}{\lambda}$.